



Investigating the influence of relevant pharmacogenetic variants on the pharmacokinetics and pharmacodynamics of orally administered docetaxel combined with ritonavir

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Abstract

The anticancer drug docetaxel exhibits large interpatient pharmacokinetic and pharmacodynamic variability. In this study, we aimed to assess the functional significance of 14 polymorphisms in the CYP3A, CYP1B1, ABCB1, ABCC2, and SLCO1B3 genes for the pharmacokinetics and pharmacodynamics of oral docetaxel, co-administered with ritonavir. None of the tested CYP3A, ABCB1, ABCC2, and SLCO1B3 genotypes and diplotypes showed a significant relation with an altered bioavailability or clearance of either docetaxel or ritonavir. Similarly, no clear effect of CYP1B1 genotype on clinical outcomes was observed in a subgroup of non-small cell lung cancer (NSCLC) patients. Our post hoc power analysis indicated that our pharmacogenetic–pharmacokinetic analysis was only powered for relatively high effect sizes, which were to be expected given the high interpatient variability. This makes it unlikely that future studies will explain the high observed interpatient variability in oral docetaxel pharmacokinetics as a result of any of these separate polymorphisms and diplotypes.

Introduction

Docetaxel is a widely used anticancer agent that triggers cell death by stabilizing microtubules. It has shown to be effective in the treatment of breast, prostate, gastric, head and neck, and non-small cell lung cancer (NSCLC) [1]. In

an attempt to improve patient burden and treatment related costs, an oral formulation of docetaxel, co-administered with the boosting agent ritonavir, has been developed. However, as with intravenous (IV) administration, oral administration of docetaxel has shown wide interpatient variability in systemic exposure which may ultimately lead to variation in clinical outcomes [2, 3].

Docetaxel is principally metabolized in hepatocytes by cytochrome P450 (CYP) 3A isoenzymes into inactive metabolites [3]. Uptake of docetaxel from the systemic circulation is mainly facilitated by the solute carriers OATP1B1 and OATP1B3 expressed on the basolateral side of hepatocytes, although other studies have also shown a role for the OATP1A2 transporter [4–6]. In addition, docetaxel and its metabolites are a substrate for the membrane-localized ATP-binding cassette (ABC) efflux transporters p-glycoprotein (ABCB1) and MRP2 (ABCC2) expressed in the apical side of hepatocytes and the epithelium of the small intestine [7, 8]. Ritonavir is similarly metabolized by CYP3A while it is also a substrate for CYP2D6. However, its role as a substrate for MRP2 and ABCB1 is not well established [9–11]. The genes encoding these enzymes and transporters are polymorphic and several of these single-nucleotide polymorphisms (SNPs) have been associated with altered functionality regarding the pharmacokinetics of

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Table 1 Allele frequencies of polymorphisms in 112 patients treated with oral docetaxel and ritonavir.

Gene	Polymorphism	Allele	Rs number	Effect	Function ^a	Allele frequency ^b		Literature q^c (Caucasian)
						<i>p</i>	<i>q</i>	
CYP3A4	−392A>G	*1B	2740574	–	increased	0.964	0.036	0.035
CYP3A4	15389C>T	*22	35599367	possible splicing defect	decreased	0.964	0.036	0.052
CYP3A7	−232A>C	*1C	45446698	–	increased	0.955	0.045	0.035
CYP3A5	6986A>G	*3C	776746	frameshift	decreased	0.058	0.942	0.930
CYP1B1	4326C>G	*3	1056836	Leu432Val	unknown	0.585	0.415	0.443
SLCO1B3	334T>G	*2	4149117	Ser112Ala	unknown	0.161	0.839	0.858
SLCO1B3	699G>A	*4	7311358	Met233Ile	unknown	0.161	0.839	0.859
SLCO1B3	767G>C	*5	60140950	Gly256Ala	unknown	0.866	0.134	0.172
ABCB1	1236C>T	*8	1128503	Gly411Gly	unknown	0.464	0.536	0.427
ABCB1	2677G>T/A	*7	2032582	Ala893Ser/Thr	unknown	0.451	0.531/0.018	0.445/0.0011
ABCB1	3435C>T	*6	1045642	Ile1145Ile	unknown	0.384	0.616	0.524
ABCC2	−24C>T	–	717620	–	unknown	0.799	0.201	0.198
ABCC2	1249G>A	–	2273697	Val417Ile	unknown	0.808	0.192	0.202
ABCC2	3972C>T	–	3740066	Ile1324Ile	unknown	0.638	0.362	0.362

^aExpected functionality of the variant protein compared to the reference protein.

^bHardy-Weinberg notation for allele frequencies with *p* the frequency for the reference allele and *q* for the variant allele.

^cVariant allele frequency as reported in dbSNP [27].

docetaxel [3]. For example; carriers of the 1236C>T (*8) polymorphism in the ABCB1 gene were found to have a significantly decreased clearance of docetaxel [12]. While patients carrying both the CYP3A4*1B and CYP3A5*1A alleles, or the CYP3A4/5*2 haplotype, showed a 64% higher clearance compared to noncarriers [13].

Furthermore, the 4326C>G (*3) polymorphism in the CYP1B1 gene has been associated with inferior survival and response to taxanes in patients carrying the variant genotype [14–16]. CYP1B1 is expressed in some tumors and while the taxanes are not among its substrates, it has been suggested that the protein encoded by the CYP1B1*3 allele reduces the efficacy of taxane treatment [17–19]. Possibly, this decreased efficacy is caused by the increased production of estradiol-3,4-quinone, an oxidation product of 4-hydroxyestradiol, which inhibits tubulin polymerization and may additionally covalently bind to docetaxel to form the inactive 4-OHE2-docetaxel adduct, thereby reducing potency [15].

To investigate whether genetic differences influence the pharmacokinetics and pharmacodynamics of oral docetaxel co-administered with ritonavir, genotyping of patients in two phase I clinical trials has been performed as a secondary endpoint [2, 20]. In this analysis, we aimed to elucidate the relationship between polymorphisms in the candidate genes CYP3A, ABCB1, ABCC2, and SLCO1B3 and the pharmacokinetics of oral docetaxel and ritonavir using a pharmacometric approach. Moreover, we investigated the relation between CYP1B1 genotype and the clinical efficacy of oral docetaxel co-administered with ritonavir.

Lastly, we performed a power analysis of two simulation scenarios in order to gain perspective on the power of the performed population pharmacogenetic study.

Methods

Clinical studies and genotyping

Docetaxel, either as a drinking solution, capsule, or tablet formulation, and co-administered with ritonavir was given to patients with various tumor types in two dose-finding phase I clinical trials (NCT01173913, ISCRTN32770468) [2, 20]. These trials investigated either a weekly once daily or a weekly twice daily schedule of administration. Both study protocols were approved by the Medical Ethics Committee of the Netherlands Cancer Institute (Amsterdam, the Netherlands). All patients had provided written informed consent prior to the start of treatment. Patients included in these studies were predominantly of Caucasian descent. We selected a panel of 14 SNPs with potential mechanistic relevance for the pharmacokinetics and/or pharmacodynamics of either ritonavir or docetaxel from literature (Table 1). In total, peripheral blood samples from 124 patients were collected in a 4 mL ethylenediaminetetraacetic acid tube and stored at −20 °C. DNA was extracted using the QIAamp DNA Mini Kit (Qiagen, Hilden, Germany) according to manufacturer's protocol and subsequently stored at −20 °C. DNA concentrations were measured at 260 nm using a NanoDrop ND-1000 UV–Vis

spectrophotometer (Thermo Fisher Scientific, Ashville, NC, USA).

Genotyping was performed using TaqMan SNP Genotyping Assays (Applied Biosystems, Carlsbad, CA, USA) according to manufacturer's protocol. The TaqMan Assay IDs and probe context sequences are shown in Supplementary Table S1. Reactions were performed using the Applied Biosystems StepOne. Quality control of the TaqMan assays was incorporated by re-measurement of two positive control samples per plate in addition to two negative controls. Positive control samples for each SNP were previously measured heterozygous and homozygous samples. These were blindly selected from a list based on their genotyping outcome from a previously performed Taqman assay. Distilled water (B. Braun Melsungen AG, Germany) was used for the negative controls.

Genotype frequencies were calculated for all polymorphisms and their distribution was tested for deviations from Hardy–Weinberg Equilibrium using the chi-square test. We performed pairwise linkage disequilibrium analysis for SNPs located on the same chromosome. In the case of significant linkage between polymorphisms an analysis of observed diplotypes could be performed.

Pharmacokinetic model

Previously we have developed an integrated semi-physiological pharmacokinetic model for oral docetaxel and oral ritonavir using nonlinear mixed effects modeling. This integrated model consists of separate pharmacokinetic models for ritonavir and docetaxel and incorporates the inhibition of docetaxel clearance, and enhancement of oral docetaxel bioavailability, by ritonavir. It was built using pharmacokinetic data from patients treated with both IV and oral docetaxel co-administered with ritonavir in several phase I studies, including the phase I studies used in the current pharmacogenetic analysis [21].

Population pharmacokinetic-pharmacogenetic analysis

To screen for potential covariates we used a systematic, two-step approach. Initially we established a list of relevant and plausible parameter-covariate relationships based on the pharmacokinetic models of oral docetaxel and ritonavir and previous knowledge of the impact of the polymorphism on protein function. Table 2 shows the parameter–covariate relationships that were considered relevant. We assumed included covariates to have an impact on the bioavailability and clearance of either docetaxel or docetaxel and ritonavir.

For 12 of the 124 genotyped patients, no corresponding pharmacokinetic data was available. We therefore matched the established genotypes with the corresponding

Table 2 Overview of potential parameter–covariate relationships.

Candidate polymorphism /Covariate	Observed in population (<i>n</i> = 112)	PK parameter of interest	Direct relation with clinical efficacy of docetaxel treatment	
CYP3A4 *1B	Wild type	104	CL _{RTV} , F _{RTV} , CL _{DOC} , F _{DOC}	None
	Heterozygous	8		
	Homozygous	0		
CYP3A4 *22	Wild type	104		
	Heterozygous	8		
	Homozygous	0		
CYP3A7 *1C	Wild type	102	CL _{RTV} , F _{RTV} , CL _{DOC} , F _{DOC}	None
	Heterozygous	10		
	Homozygous	0		
CYP3A5 *3C	Wild type	2	CL _{RTV} , F _{RTV} , CL _{DOC} , F _{DOC}	None
	Heterozygous	9		
	Homozygous	101		
CYP1B1 *3	Wild type	37	none	ToS, Response
	Heterozygous	57		
	Homozygous	18		
SLCO1B3 *2/ *4 diplotype	TG/TG	0	CL _{DOC} , F _{DOC}	None
	Non-GA/GA- TG/TG	36		
	GA/GA	76		
SLCO1B3 *5	Wild type	83		
	Heterozygous	28		
	Homozygous	1		
ABCB1 *8	Wild type	16	CL _{DOC} , F _{DOC}	None
	Heterozygous	72		
	Homozygous	24		
ABCB1 *7	Wild type	13		
	Heterozygous	73		
	Homozygous	22		
	Heterozygous	2		
	Heterozygous	2		
ABCB1 *6	Wild type	15		
	Heterozygous	56		
	Homozygous	41		
ABCB1 diplotype	CGC/CGC	8	CL _{DOC} , F _{DOC}	None
	Non-CGC/ CGC-TTT/ TTT	91		
	TTT/TTT	25		
ABCC2 –24C>T	Wild type	70	CL _{DOC} , F _{DOC}	None
	Heterozygous	39		
	Homozygous	3		
ABCC2 1249G>A	Wild type	73		
	Heterozygous	35		
	Homozygous	4		
ABCC2 3972C>T	Wild type	46		
	Heterozygous	51		
	Homozygous	15		

PK pharmacokinetic, CL clearance, F bioavailability, ToS time on study, DOC docetaxel, RTV ritonavir.

pharmacokinetic data for 112 out of 173 patients in our population pharmacokinetic model. Since the availability of genotyping data was dependent on the clinical studies in which patients were included, we assumed that this data was

missing completely at random for the patients who were not genotyped ($n = 61$). Missing data were handled by estimating an extra parameter for the effect of a missing covariate in this population, as described previously [22]. All model estimations were repeated after the inclusion of this extra parameter.

Subsequently, the distribution of individual Bayesian estimates of the pharmacokinetic parameter of interest (eta (η) distribution) versus the genotype were visually explored. These were related to the population parameter estimates according to either Eq. (1) or Eq. (2).

$$P_i = P \cdot \exp(\eta_{i,BSV}) \quad (1)$$

$$P_i = P \cdot \exp(\eta_{i,BSV} + \eta_{i,BOV}) \quad (2)$$

With P_i the individual parameter estimate for individual i , P the population parameter estimate, and η_i either between-subject variability or between-occasion variability effect distributed following $N(0, \omega^2)$. Potentially relevant genotypes were added to the list of parameter–covariate relationships to be formally tested.

We used an automated stepwise covariate modeling (SCM) algorithm to perform this population pharmacokinetic–pharmacogenetic analysis [23, 24]. In SCM, each potential parameter–covariate relation as defined in the first step was tested univariately on the original pharmacokinetic model for a significant improvement of the model fit to the data. As quantified by the difference in objective function value (dOFV). Statistical evaluation was performed using the log-likelihood ratio test. For forward inclusion and backward elimination significance levels of $p < 0.05$ (degrees of freedom (df) = 1, dOFV = -3.84 ; df = 2, dOFV = -5.991) and $p < 0.001$ (df = 1, dOFV = -10.83 ; df = 2, dOFV = -13.82) were used, respectively. The effect of different genotypes on parameter P was coded proportionally to the respective pharmacokinetic parameter. For example,

$$P = P_{pop} \times (1 + \theta_1^{\text{heterozygous}} \times \theta_2^{\text{homozygous}}) \quad (3)$$

where P_{pop} is the population parameter estimate for the wild type, and *homozygous* and *heterozygous* can be either 0 or 1, therefore θ_1 and θ_2 are the separate fixed effects for either the heterozygous or the homozygous variant genotype. For genotypes with low minor allele frequencies resulting in <3 patients with a variant genotype, such as CYP3A5*1/*1 ($n = 2$), the homozygous covariate effect was assumed to be two times the effect of the heterozygous genotype (Eq. (4)) [12], provided that this was mechanistically plausible and correlated with the

observed differences in interpatient variability.

$$P = P_{pop} \times (1 + \theta_3 \times \text{genotype}) \quad (4)$$

With θ_3 the effect of the variant allele and genotype either 0 (wild-type), 1 (heterozygous), or 2 (homozygous variant).

Genotype related to time on study and response

We performed a separate pharmacodynamic analysis in which we investigated the relationship between genotype and clinical efficacy of oral docetaxel treatment. The majority of patients in both phase I trials were diagnosed with NSCLC. In order to attain a substantial population with similar clinical prognosis our pharmacodynamic analysis was performed only on this subgroup. SNPs which were thought to have relevant pharmacogenetic–pharmacodynamic relations are displayed in Table 2. Clinical outcome of patients by genotype was visualized using swimmer plots displaying time on study and response to treatment for each individual patient. Time on study was defined as the time in months from start of treatment until either progression or cessation of treatment due to other causes. Tumor response was evaluated in accordance with RECIST version 1.0 [25].

Stochastic simulation and estimation with parametric power estimation

We conducted trial simulations and power estimations for two different hypothetical scenarios using stochastic simulation and estimation in order to aid the interpretation of the results of our pharmacogenetic–pharmacokinetic analysis. For this, we simulated a data frame that included a virtual polymorphism as a covariate. The data frame was similar to the dataset used in SCM. It contained pharmacokinetic data from 173 patients of which 112 were genotyped. The assumption was made that the allele frequency of this variant allele in the population was 0.45, which results in 25% (28/112 patients) of the population having the homozygous variant genotype. For both scenarios we assumed only an effect of the homozygous variant genotype on docetaxel clearance. For the first scenario we created a full pharmacokinetic model in which carriers of this homozygous variant genotype had a fixed 25% decreased clearance of docetaxel ($\theta_{\text{virtualSNP}} = -0.25$). In the second scenario, we assumed the effect of the homozygous variant genotype to be 40% ($\theta_{\text{virtualSNP}} = -0.40$). Similar to our covariate analysis, missing genotype was assumed to be missing completely at random in this simulation study. Thus, for the group of patients for whom genotype was missing ($n = 61$), the SNP effect was

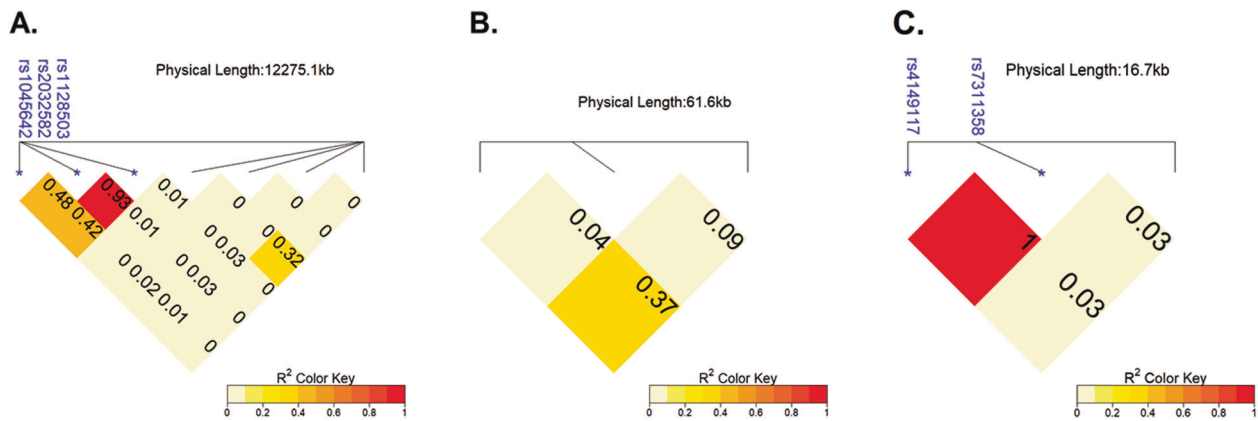


Fig. 1 Linkage disequilibrium plots. **A.** Heatmap of pairwise linkage disequilibrium for SNPs in the CYP3A and ABCB1 genes on chromosome 7, **B.** ABCB2 polymorphisms on chromosome 10, and **C.** SLCO1B3 polymorphisms on chromosome 12.

imputed as follows:

$$P = P_{pop} \times (1 + (\theta_{virtualSNP} \times f_{prevalence})) \quad (5)$$

With $\theta_{virtualSNP}$ the effect of the polymorphism on the population parameter estimate, and f the prevalence of the polymorphism. We created a nested reduced version of both models by fixing the effect of the covariate to zero. The full model and simulated data frame were then used to stochastically simulate 100 datasets which were subsequently used in the parameter estimation of both the full and reduced model. A previously described parametric power estimation algorithm (PPE) was used to generate a power versus study size curve [26].

Software

Model estimations and simulations were performed using nonlinear mixed effects modeling software (NONMEM, version 7.3, ICON Development Solutions, Ellicott City, MD, USA) together with a gfortran compiler. Pirana (version 2.9.9, Certara, Princeton, New Jersey) was used as a graphical interface. The model estimations used first-order conditional estimation with interaction. We used the SCM and SSE tools from the Perl-speaks-NONMEM toolkit (PsN, version 4.7.0) [24, 26]. R (version 3.6.2) and RStudio (version 1.2.5042; Boston, MA, USA) were used for data processing, genetic analysis, plotting, and implementation of the PPE algorithm.

Results

Genotyping

Minor allele frequencies for the SNPs analyzed corresponded well with those observed in the Caucasian

population as shown in Table 1 [27]. Average call rates were >99% for each SNP. The frequency distributions of most SNPs were in Hardy-Weinberg equilibrium. However, deviations from Hardy-Weinberg equilibrium were observed for the ABCB1*7 ($p = 0.0027$) and ABCB1*8 ($p = 0.002$) polymorphisms (calculations not shown). For ABCB1, linkage of the SNPs at the 1236, 2677, and 3435 loci was observed (Fig. 1A) which is in line with previous reports [28]. And as expected, we also observed a strong genetic linkage of the SLCO1B3*2 and SLCO1B3*4 SNPs ($r^2 = 1$) (Fig. 1C) [29]. Due to the complete linkage of the SLCO1B3*2 and *4, these SNPs were only analyzed as a diplotype. For the ABCB1 SNPs an analysis of the diplotype was performed in addition to the analysis of the separate SNPs.

Population pharmacokinetic–pharmacogenetic analysis

The distributions of the between-subject variability for each patient (η) (Supplementary Figs. S1 and S2.) were visually examined and formed the basis for the selection of the parameter–covariate relationships to be investigated using SCM.

The effects of including the polymorphisms and diplotypes as covariates on both pharmacokinetic models and the corresponding effect sizes are shown in Table 3. We identified no statistically significant associations between variants of the candidate genes and any of the relevant pharmacokinetic parameters in either the oral docetaxel or the ritonavir model. Furthermore, the estimated (non-significant) effect sizes were relatively low in comparison with the clinically observed inter-patient variability in pharmacokinetics and had a high imprecision.

Table 3 Estimates of the effect of inclusion of relevant covariates in the docetaxel and ritonavir pharmacokinetic models.

PK parameter	Parameter estimate (RSE (%))	Unit	Covariate	df	Genotype	Covariate effect size (RSE (%))	P value	Missing parameter effect size (RSE (%))
Ritonavir model								
Clearance (CL_{RTV})	8.27 (10)	L/h	CYP3A4 *1B ^b	1	Heterozygous	0.18	0.49	1.25 (17)
			CYP3A4 *22 ^b	1	Heterozygous	-0.18	0.39	
Population relative gut bioavailability of ritonavir after first dose ($F_{RTV, intrinsic}$)	1 ^a	-	CYP3A5 *3C ^c	1	homozygous wild-type/ heterozygous	-0.02	0.93	1.49 (19)
Docetaxel model								
Population uninhibited intrinsic clearance (CL_{int0})	1980 ^a	L/h	ABCC2 3972C>T	2	Heterozygous	0.03 (492)	0.69	1.02 (13)
					Homozygous variant	-0.11 (127)		
			CYP3A7 *1C ^b	1	Heterozygous	-0.26 [60]	0.08	
			ABCB1 *8	2	Heterozygous	0.10 (194)	0.63	
					Homozygous variant	-0.01 (2462)		
			ABCB1 *7 ^c	2	Heterozygous (GT/GA)	0.08 (240)	0.29	
					TT/TA genotypes	-0.10 (142)		
			SLCO1B3 *2/*4 diplotype ^b	1	non-GA/GA-TG/TG	-0.20 (46)	0.04	
Population gut bioavailability in combination with ritonavir relative to without ($F_{ritonavir}$)	3.75 (28)	-	CYP3A4 *1B ^b	1	Heterozygous	-0.31 (60)	0.06	0.92 (12)
			CYP3A4 *22 ^b	1	Heterozygous	0.18 (187)	0.37	
			ABCC2 1249G>A	2	Heterozygous	0.18 (70)	0.06	
					Homozygous	-0.34 (67)		
			ABCC2 -24C>T	2	Heterozygous	-0.08 (112)	0.16	
					Homozygous	0.63 (65)		
			CYP3A5 *3 ^c	1	Combined homozygous wild-type/heterozygous	-0.04 (703)	0.72	
			ABCB1 diplotype	2	Non-CGC/CGC-TTT/TTT	0.18 (182)	0.46	
					TTT/TTT	0.29 (119)		

PK pharmacokinetic, *df* degrees of freedom, RSE relative standard error

^aFixed population parameter estimate.

^bLow minor allele frequency, resulting in the absence of a homozygous variant/wild type genotype in this population.

^cLow minor allele frequency, effect of the covariate therefore estimated using Eq. (4).

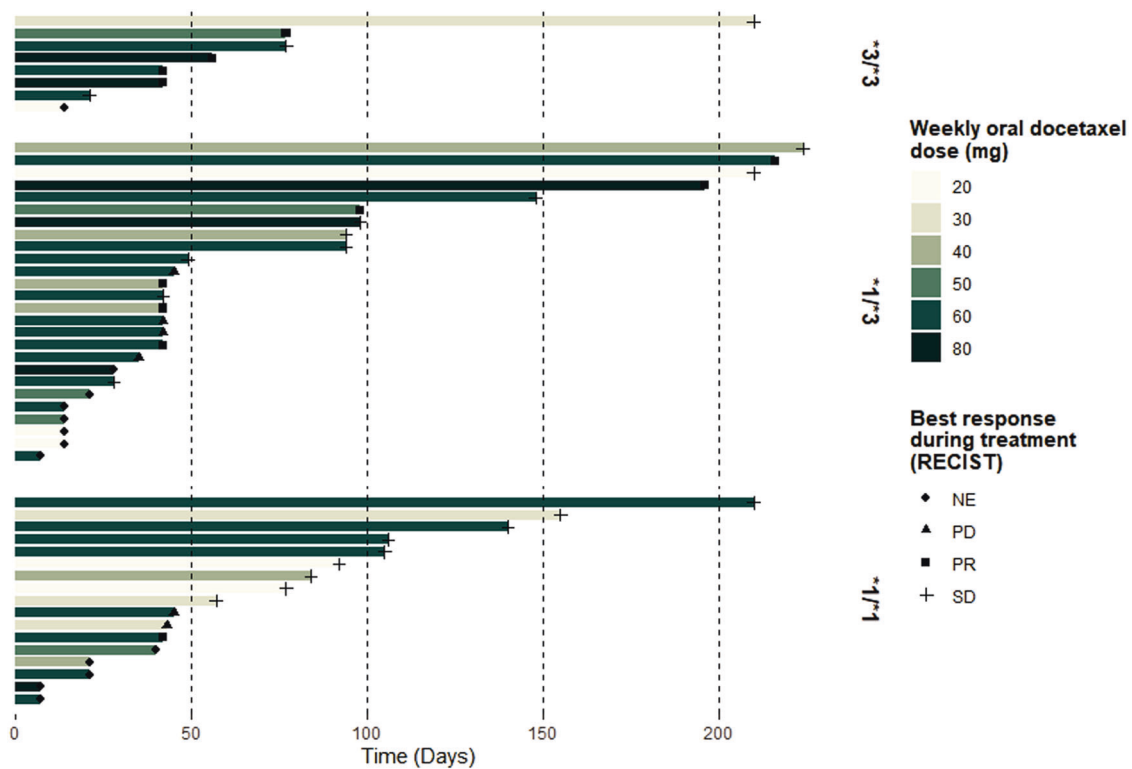


Fig. 2 Swimmer plot visualizing time on study and best outcome of treatment with oral docetaxel for 51 NSCLC patients grouped by CYP1B1 genotype. NE not evaluable; PD, progressive disease; PR partial response, SD stable disease. *1, wild-type allele; *3, variant allele.

Genotype related to time on study and response

We identified a subgroup of 51 patients with locally advanced or metastatic NSCLC from both phase I trials for which corresponding genotype data was available. We visualized the time on study and outcome per patient, received daily dose, and CYP1B1 genotype using a swimmers plot as shown in Fig. 2. We observed no distinct differences in the time on study between patients with the *3/*3 genotype and patients with either the *1/*3 and *1/*1 genotype. We did observe a higher objective response rate in patients with the *3/*3 genotype compared to the *1/*3 and *1/*1 genotypes (50% vs 16%) while the rate of patients who achieved stable disease as best response to treatment was higher in the *1/*3 and *1/*1 genotypes (49% vs 38%). The median weekly dose administered in both groups was 60 mg.

Stochastic simulation and estimation with parametric power estimation

The power versus study size for the two scenarios is shown in Fig. 3. Figure 3A illustrates that using our current analysis, a sample size of 220 patients would be required to achieve 80% power to detect a significant 25% decreased clearance of oral docetaxel as a result of an SNP with a

prevalence of 25%. To achieve 80% power to detect an effect of a polymorphism that causes a 40% decrease in clearance and a prevalence of 25%, a sample size of 85 patients would be required. This indicates that our current analysis would have been sufficiently powered in the latter scenario.

Discussion

We performed a pharmacogenetic analysis of the effect of several polymorphisms in genes generally considered relevant for the pharmacokinetics and pharmacodynamics of docetaxel and ritonavir in a cohort of cancer patients, in an attempt to unravel the causes for differences in outcomes and the extensive interpatient variability in exposure to oral docetaxel.

This is not the first study investigating the influence of genetic differences on the pharmacokinetics of docetaxel [3]. We are, however, the first to examine the influence of genetic polymorphisms in a population of patients treated with oral docetaxel formulations in combination with the boosting agent ritonavir. From a pharmacokinetic point of view, orally administered docetaxel is subject to more interactions with these enzymes and transporters due to its uptake from the gastrointestinal tract and subsequent first-pass metabolism [21].

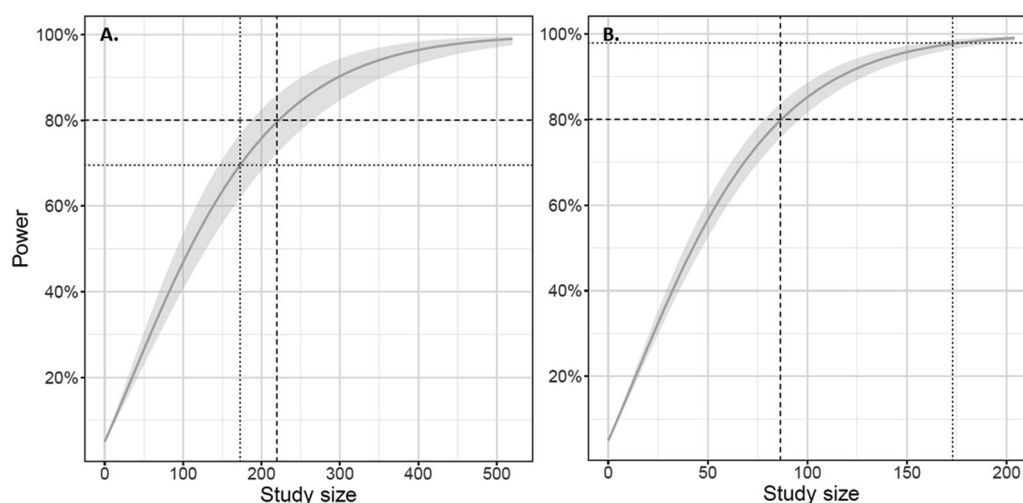


Fig. 3 PPE power curves. Curves were generated for a full model (containing either a covariate effect $\theta_{\text{virtualSNP}} = -0.25$ (A.) or a covariate effect of $\theta_{\text{virtualSNP}} = -0.40$ (B.)), both with a prevalence of 25%, versus a reduced model estimated on 100 stochastically simulated datasets. The PPE curves indicate that a study size of ~220 patients would be required to achieve 80% power in the case of an

effect size of -0.25 whilst with an effect size of -0.40 a population of ~85 patients would be required to detect a significant effect of this polymorphism. Dashed line, number of patients to achieve 80% power threshold; dotted line, actual power using a study size of 173 patients; shaded region, uncertainty in the power estimate.

Our population displayed variant allele frequencies that corresponded reasonably with those observed in the Caucasian population. However, some of the allele frequencies of the investigated SNPs showed deviations from Hardy–Weinberg equilibrium. Most likely this observation is caused by genetic drift as a result of the limited number of patients included in this analysis. As expected, we observed strong linkage of the 334T>G and 699G>A polymorphisms in the *SLCO1B3* gene and polymorphisms in the *ABCB1* gene at positions 1236, 2677, and 3435 [28, 29]. Conversely, no linkage was observed for the *CYP3A4**1B and *CYP3A5**1 polymorphisms, that are known to bring about a known haplotype in the *CYP3A* gene [13, 30].

We did not identify a significant effect of any of the polymorphisms and diplotypes in the *CYP3A*, *ABCB1*, *ABCC2*, and *SLCO1B3* genes on the pharmacokinetics of oral docetaxel and ritonavir.

Nevertheless, some of the investigated polymorphisms appeared to show mechanistically plausible trends towards significance when univariately tested as a covariate to the pharmacokinetic model for oral docetaxel. Patients who were heterozygous carriers of the *CYP3A4**1B polymorphism displayed a trend towards a decreased bioavailability of docetaxel (31% decrease, $p = 0.06$). A trend that was also observed for homozygous carriers of the *ABCC2* 1249G>A polymorphism (34% decrease, $p = 0.06$). Both these effects are in accordance with the previously reported effect of these SNPs on taxane pharmacokinetics [31, 32]. Despite not resulting in a significant improvement of the models, the *CYP3A4**22 polymorphism displayed estimated effect sizes consistent with the previously reported functionality (21%

reduced *CYP3A4* activity) of the polymorphism when added to the ritonavir (18% decreased clearance) and docetaxel model (18% increased bioavailability) [33].

In part, the absence of an observed effect may be explained by the co-administration of ritonavir. Ritonavir is a strong and irreversible inhibitor of *CYP3A4* ($IC_{50} = 0.014 \mu\text{M}$) and *CYP3A5* ($IC_{50} = 0.09 \mu\text{M}$) and a weak inhibitor of *CYP3A7* ($IC_{50} = 0.6 \mu\text{M}$) [34, 35]. Additionally, it is a strong inhibitor of *ABCB1* ($IC_{50} = 0.2 \mu\text{M}$) and a general OATP inhibitor ($IC_{50} = 1.3\text{--}6.1 \mu\text{M}$) [36, 37]. While some studies also suggest ritonavir may inhibit *ABCC2* [38]. As a result, inherited differences in the functionality of these enzymes and transporters may have been masked in this study due to strong inhibition. Moreover, the effect of deficiencies in for example the *OATP1B3* transporter may have been indiscernible due to compensatory mechanisms such as *OAT2* for which docetaxel is also a substrate [39].

We also investigated the potential relation between polymorphisms in the *CYP1B1* gene and clinical outcomes of oral docetaxel treatment. We selected a group of NSCLC patients who composed the largest subgroup of the various tumor types included in the original phase I trials. As opposed to previous studies we observed no detrimental effect of the *CYP1B1**3/*3 genotype on clinical outcomes after treatment with oral docetaxel [14–16]. Contrary to our expectations, we even observed a higher response rate in patients with the *3/*3 genotype as opposed to the *1/*3 and *1/*1 genotype.

The previously reported effect of this variant polymorphism was more pronounced in two clinical trials in

prostate cancer patients [15, 16] as compared to the trial in NSCLC patients [14], which may suggest a more dominant role of the CYP1B1*3 polymorphism in the former tumor type.

We should note that the dismal prognosis of NSCLC patients after first line therapy, in addition to the large variation in administered treatments and limited sample size, may have obscured any effect of the CYP1B1 genotype and complicate the drawing of firm conclusions based on our analysis.

We performed a post hoc power study to aid the interpretation of the results of our pharmacokinetic–pharmacogenetic analysis and ultimately to investigate the feasibility of studying these polymorphisms in larger cohorts. We reasoned that in order to be clinically relevant, the effect of a polymorphism would have to be sufficiently large to explain the aforementioned variability in docetaxel pharmacokinetics [2, 20]. We therefore simulated two separate scenarios incorporating a virtual polymorphism with either an effect of a 25% or a 40% decreased docetaxel clearance and both with a clinically relevant prevalence of 25%.

The first simulation scenario demonstrated that a study size of about 220 patients would be required in order to achieve 80% power, indicating that our current analysis was not sufficiently powered in the event of a polymorphism with such an effect size and prevalence. In relation to the observed effect sizes in our own pharmacokinetic–pharmacogenetic analysis (Table 3) we can see that some of the polymorphisms indeed exhibit effect sizes with a similar magnitude. With the second scenario we sought to demonstrate the required study size to achieve 80% power in the event of a highly clinically relevant polymorphism with an effect size of 40%. In this case a study size of 85 patients would result in sufficient power. Meaning our study would have been sufficiently powered to detect these larger differences in pharmacokinetics as a result of these polymorphisms.

We should note that in both simulation studies we handled the missing covariate data for 61 patients by imputing the population mean of the effect of the covariate times the prevalence. Thereby using a study set up similar to our original pharmacokinetic–pharmacogenetic analysis and assuming that the missingness of the data occurred completely at random.

In conclusion, we found no significant effect of the candidate CYP3A, CYP1B1, ABCB1, ABCC2, and SLCO1B3 polymorphisms and diplotypes on the pharmacokinetics and pharmacodynamics of oral docetaxel and ritonavir in our analysis of clinical phase I data. A post hoc power analysis indicated that our pharmacokinetic–pharmacogenetic analysis was probably underpowered, given the encountered polymorphism prevalences and moderate effect sizes. Thus, based on our current study, we cannot discard the hypothesis that

polymorphisms in the investigated genes contribute to some extent to the clinically observed interpatient variability in oral docetaxel pharmacokinetics. However, it can be concluded that it is unlikely that future studies will explain the high observed interpatient variability in oral docetaxel pharmacokinetics as a result of any of these separate polymorphisms and diplotypes. This finding may support feasibility assessment of larger prospective clinical trials investigating the influence of genetic polymorphisms on the pharmacokinetics of oral docetaxel combined with ritonavir.

Code availability

The NONMEM control streams for the oral docetaxel and ritonavir models that were used in our pharmacokinetic–pharmacogenetic analysis and in our trial simulations are available in the supplementary material.

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Compliance with ethical standards

Conflict of interest JHB holds a patent on oral taxane formulations and is a (part-time) employee and (indirect) shareholder of Modra Pharmaceuticals, a spin-off company that develops oral taxane formulations. The other authors declare that they have no conflict of interest.

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